

## SOME REMARKS ON THE LARGE SIEVE OF YU. V. LINNIK

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### § 1. Introduction

YU. V. LINNIK has discovered (see [1]) in 1941 a very powerful new method of elementary number theory, which he called the large sieve<sup>1</sup>. In his original formulation the large sieve asserts that if we take any sequence  $S_N$  consisting of  $Z$  positive integers  $\leq N$ , and if  $Y$  denotes the number of those primes<sup>2</sup>  $p \leq \sqrt{N}$  for which all the elements of the sequence  $S_N$  are contained in  $\leq p(1 - \varepsilon)$  residue classes mod  $p$ , where  $0 < \varepsilon < 1$ , then one has

$$(1.1) \quad Y \leq \frac{20\pi N}{\varepsilon^2 Z}.$$

As shown by the second named author in [3], Linnik's method is capable to prove much more, namely that if  $Z$  is not too small compared with  $N$ , then the elements of the sequence  $S_N$  not only occupy "almost all" residue classes mod  $p$  with respect to most primes  $p \leq \sqrt{N}$ , but are almost uniformly distributed in the  $p$  residue classes mod  $p$  for most primes  $p \leq \sqrt{N}$ . More exactly, let us denote by  $Z(a, p)$  (where  $a = 0, 1, \dots, p-1$ ) the number of elements of the sequence  $S_N$  which are congruent to  $a$  mod  $p$ . Then one has, putting

$$(1.2) \quad \Delta^2(p) = p \sum_{a=0}^{p-1} \left( Z(a, p) - \frac{Z}{p} \right)^2.$$

the inequality<sup>3</sup>

<sup>1</sup> As regards important applications of the large sieve in number theory, see e.g. [2] 3], [4], [5], [6]; [5] and [6] contain many further references.

<sup>2</sup> In this paper  $p$  always denotes a prime number.

<sup>3</sup> Here and in what follows all the constants of the  $O$ -estimates are absolute, i.e. do not depend on  $N$ , nor on the sequence  $S_N$  nor on  $Q$ .

$$(1.3) \quad \sum_{p \leq Q} \Delta^2(p) = O(Z^{2/3} N^{4/3} Q^{1/3})$$

for  $Q \leq N^{3/5}$ . Later, the second named author has found (see [7]), a new probabilistic method for proving theorems of the type of the large sieve. This method (developed further and generalized in the papers [8], [9], [10], [11], [12]) gave the result

$$(1.4) \quad \sum_{p \leq Q} \Delta^2(p) = O(Z(Q^3 + N))$$

for  $Q \leq \sqrt{N}$ . This estimate is better than (1.3) for  $Q \leq N^{3/8}$ , but weaker if  $N^{3/8} < Q \leq \sqrt{N}$ .

Especially for  $Q = N^{1/3}$  this result gives

$$(1.5) \quad \sum_{p \leq N^{1/3}} \Delta^2(p) = O(NZ).$$

The estimate (1.5) is essentially best possible, because if for instance  $S_N$  is the sequence of odd numbers  $\leq N$ , one has  $Z(0, 2) = 0$  and thus  $2 \left( Z(0, 2) - \frac{Z}{2} \right)^2 = \frac{Z^2}{2}$  i.e. this single term is already of order  $NZ$ .

The probabilistic approach, besides leading to a very sharp estimate for  $Q \leq N^{1/3}$ , has thrown light on the reasons why an arbitrary sufficiently dense subsequence of the sequence  $1, 2, \dots, N$  has to be almost uniformly distributed among the residue classes mod  $p$  for most  $p \leq N^{1/3}$ ; it became obvious that this is due to the statistical independence (more exactly: almost independence) of the distribution mod  $p$  and mod  $q$  of the numbers  $n \leq N$  for any two primes  $p, q \leq N^{1/3}$ ,  $p \neq q$ .

In the last two years important progress was made on the large sieve. The first essential improvement was obtained by K. F. ROTH [13]. His result was sharpened by BOMBIERI [14] who has shown that (1.5) holds also for  $Q = \sqrt{N}$ . More exactly Bombieri proved

$$(1.6) \quad \sum_{p \leq Q} \Delta^2(p) = O(Z(Q^2 + N)).$$

Clearly (1.6) is superior to both (1.3) and (1.4) for the full ranges  $Q \leq N^{3/5}$  resp.  $Q \leq \sqrt{N}$ .

An important generalization of Bombieri's theorem has been obtained by H. DAVENPORT and H. HALBERSTAM [15]. To make this advance clear one has to notice that putting

$$(1.7) \quad S(x) = \sum_{n \in S_N} e^{2\pi i n x}$$

one has

$$(1.8) \quad \Delta^2(p) = \sum_{a=1}^{p-1} \left| S\left(\frac{a}{p}\right) \right|^2.$$

Now DAVENPORT and HALBERSTAM have proved that if  $\alpha_1, \alpha_2, \dots, \alpha_D$  are arbitrary real numbers in the interval  $(0, 1)$  such that  $|\alpha_i - \alpha_j| \cong \delta > 0$  for  $i \neq j$ , one has

$$(1.9) \quad \sum_{i=1}^D |S(\alpha_i)|^2 = O\left(Z\left(\frac{1}{\delta} + N\right)\right).$$

Clearly, if the numbers  $\frac{a}{p}$  ( $a = 1, 2, \dots, p-1$ ;  $p \leq Q$ ) are taken as the numbers  $\alpha_1, \dots, \alpha_D$  ( $D = \sum_{p \leq Q} (p-1)$ ) then  $\delta \cong \frac{1}{Q^2}$  and thus (in view of (1.8)) (1.9) implies (1.6).

Note that from (1.9) one obtains even more than (1.6), namely that

$$\sum_{q \leq Q} \sum_{\substack{a=1 \\ (a, q)=1}}^q \left| S\left(\frac{a}{q}\right) \right|^2 = O\left(Z\left(\frac{1}{\delta} + N\right)\right)$$

because if  $(a, q) = 1$ ,  $(a', q') = 1$  (here  $(a, q)$  denotes the greatest common divisor of  $a$  and  $q$ ) one has for  $q, q' \leq Q$  and  $\frac{a}{q} \neq \frac{a'}{q'}$

$$\left| \frac{a}{q} - \frac{a'}{q'} \right| \cong \frac{1}{Q^2}.$$

Recently P. X. GALLAGHER [16] has found a very elegant and simple method for proving (1.9). More exactly, he proved

$$(1.10) \quad \sum_{\nu=1}^D |S(\alpha_\nu)|^2 \cong Z\left(\frac{1}{\delta} + \pi N\right)$$

which implies by (1.8)

$$(1.11) \quad \sum_{p \leq Q} A^2(p) \cong Z(Q^2 + \pi N).$$

Thus we have for  $Q = \sqrt{N}$

$$(1.12) \quad \sum_{p \leq Q} A^2(p) \cong (\pi + 1)ZN.$$

In the paper [17] of the first named author it has been mentioned (without giving the proof in detail) that by a probabilistic argument it can be shown that (1.12) cannot hold if  $Q$  is of larger order of magnitude than  $\sqrt{N \log N}$ . The aim of the present paper is to prove this statement in detail, and to get some related results concerning the behaviour of  $\sum_{p \leq Q} A^2(p)$ , when  $S_N$  is a random subset of the set  $\{1, 2, \dots, N\}$ .

The results obtained throw some light on certain open problems connected with the large sieve.

## § 2. Equidistribution of random sequences in arithmetic progressions

In this § let  $S_N$  denote a random subsequence of the sequence  $\{1, 2, \dots, N\}$  obtained as follows: let  $\varepsilon_1, \varepsilon_2, \dots, \varepsilon_N$  be independent random variables, each of which takes on the values 1 and 0 with probability  $\frac{1}{2}$ ; let  $S_N$  denote the set of those  $n \leq N$  for which  $\varepsilon_n = 1$ . (It is easy to see that under these suppositions each of the  $2^N$  subsets of the set  $\{1, 2, \dots, N\}$  has the same probability to be chosen.) In this case

$$(2.1) \quad Z = \sum_{n=1}^N \varepsilon_n$$

$$(2.2) \quad Z(a, p) = \sum_{k=0}^{\left[\frac{N-a}{p}\right]} \varepsilon_{kp+a}$$

and consequently

$$(2.3) \quad \Delta^2(p) = p \cdot \sum_{a=0}^{p-1} \left( Z(a, p) - \frac{Z}{p} \right)^2$$

are all random variables. One obtains easily

$$(2.4) \quad \Delta^2(p) = p \cdot \sum_{a=0}^{p-1} Z^2(a, p) - Z^2$$

and thus, putting<sup>4</sup>

$$(2.5) \quad \pi_r(Q) = \sum_{p \leq Q} p^r \quad (r=0, 1, 2, \dots)$$

we have

$$(2.6) \quad R(Q) = \sum_{p \leq Q} \Delta^2(p) = \sum_{n=1}^N \sum_{m=1}^N [A_Q(n-m) - \pi_0(Q)] \varepsilon_n \varepsilon_m,$$

where

$$(2.7) \quad A_Q(k) = \sum_{\substack{p/k \\ p \leq Q}} p$$

and thus  $A_Q(-k) = A_Q(k)$  and especially

$$(2.8) \quad A_Q(0) = \pi_1(Q).$$

Let us determine first the expectation<sup>5</sup> of  $R(Q)$ . As

$$(2.9) \quad E(\varepsilon_n) = \frac{1}{2} \quad \text{and} \quad E(\varepsilon_n \varepsilon_m) = \begin{cases} \frac{1}{4} & \text{if } n \neq m \\ \frac{1}{2} & \text{if } n = m \end{cases}$$

<sup>4</sup> Thus  $\pi_0(Q)$  denotes the number of primes  $\leq Q$ .

<sup>5</sup> The expectation of a random variable  $\xi$  will be denoted by  $E(\xi)$ .

we obtain

$$(2.10) \quad E(R(Q)) = \frac{1}{4} \sum_{n=1}^N \sum_{m=1}^N (A_Q(n-m) - \pi_0(Q)) + \frac{N}{4} \pi_1(Q).$$

Now clearly

$$(2.11) \quad \sum_{n=1}^N \sum_{m=1}^N [A_Q(n-m) - \pi_0(Q)] = \sum_{p \leq Q} p \left\{ \sum_{a=0}^{p-1} \left[ \left[ \frac{N-a}{p} \right] + 1 \right]^2 - \frac{N^2}{p^2} \right\}.$$

Let us suppose that  $N \equiv r \pmod{p}$ , where  $0 \leq r < p$ . Then we have

$$(2.12) \quad p \cdot \left\{ \sum_{a=0}^{p-1} \left[ \left[ \frac{N-a}{p} \right] + 1 \right]^2 \right\} = 2N + r(p-r) + p - 2r.$$

Thus it follows that

$$(2.13) \quad \sum_{n=1}^N \sum_{m=1}^N [A_Q(n-m) - \pi_0(Q)] = 2N\pi_0(Q) + O(Q^2\pi_0(Q))$$

and thus, taking into account that  $\pi_0(Q) = \frac{Q}{\log Q} + O\left(\frac{Q}{\log^2 Q}\right)$  and

$$(2.14) \quad \pi_1(Q) = \frac{Q^2}{2 \log Q} + O\left(\frac{Q^2}{\log^2 Q}\right)$$

it follows

$$(2.15) \quad E(R(Q)) = \frac{NQ\pi_0(Q)}{8} + O(Q^2\pi_0(Q)) + O\left(\frac{Q^2N}{\log^2 Q}\right).$$

Thus the expectation of  $R(Q)$  is smaller by a factor of order  $\frac{1}{\log Q}$  as  $NQ^2$ .

Note that the expectation of  $R(Q)$  can be interpreted as its average over all  $2^N$  subsequences of the sequence  $\{1, 2, \dots, N\}$ . Thus the average of  $R(Q)$  is of order  $O(N^2)$  even for  $Q = O(\sqrt{N \log N})$  while for its maximum according to (1.6) this is known only for  $Q = O(\sqrt{N})$ . It is an open question whether the estimate

$$(2.16) \quad R(Q) = O(N^2)$$

holds for all sequences  $S_N$  if  $Q \sim \sqrt{N} \psi(N)$  for some function  $\psi(N)$  such that  $\psi(N) \rightarrow \infty$  for  $N \rightarrow \infty$ . Our method is not capable of giving such a result; however by evaluating the variance of the random variable  $R(Q)$  we can show by Čebishev's inequality that the estimate (2.16) is valid at least for most subsequences  $S_N$ .

To evaluate the variance<sup>6</sup> of  $R(Q)$  note that though the random variables  $\varepsilon_n \varepsilon_m$  are not independent, they are pairwise uncorrelated and thus the variance

<sup>6</sup> The variance of a random variable  $\xi$  will be denoted by  $D^2(\xi)$ .

of the sum on the right hand side of (2.6) is equal to the sum of the variances of the single terms. As  $D^2(\varepsilon_n \varepsilon_m) = \frac{3}{16}$  if  $n \neq m$  and  $D^2(\varepsilon_n^2) = \frac{4}{16}$ .

$$(2.17) \quad D^2(R(Q)) = \frac{3}{16} \sum_{n=1}^N \sum_{m=1}^N (A_Q(n-m) - \pi_0(Q))^2 + \frac{N}{16} (\pi_1(Q) - \pi_0(Q))^2.$$

Now clearly

$$(2.18) \quad \sum_{n=1}^N \sum_{m=1}^N A_Q^2(n-m) \cong N^2[\pi_0^2(Q) + \pi_1(Q) - \pi_0(Q)] + 2N\pi_1^2(Q) + \pi_2^2(Q).$$

As further from (2.11) we have

$$\sum_{n=1}^N \sum_{m=1}^N A_Q(n-m) \cong N^2\pi_0(Q) - 2N\pi_0(Q) + \pi_1(Q)$$

it follows

$$(2.19) \quad D^2(R(Q)) \cong \frac{3}{16} N^2(\pi_1(Q) - \pi_0(Q)) + \\ + \frac{3N}{8} \left( \pi_1^2(Q) + 2\pi_0^2(Q) + \frac{(\pi_1(Q) - \pi_0(Q))^2}{6} \right) + \frac{3}{16} \pi_2^2(Q) - \frac{3}{8} \pi_1(Q)\pi_0(Q).$$

In view of (2.14), it follows

$$(2.20) \quad D^2(R(Q)) = O\left(\frac{N^2Q^2}{\log Q}\right) + O\left(\frac{NQ^4}{\log^2 Q}\right)$$

i.e.

$$(2.21) \quad \frac{D(R(Q))}{E(R(Q))} = O\left(\frac{\sqrt{\log Q}}{Q}\right) + O\left(\frac{1}{\sqrt{N}}\right).$$

It follows from Čebishev's inequality that for  $\lambda > 1$  with probability  $\cong 1 - \frac{1}{\lambda^2}$   $R(Q)$  is contained in an interval

$$[E(R(Q)) - \lambda D(R(Q)), E(R(Q)) + \lambda D(R(Q))].$$

Choosing for  $\lambda$  the value  $\lambda = \min\left(\frac{Q}{(\log Q)^{3/2}}, \frac{\sqrt{N}}{\log Q}\right)$  it follows that for all but  $\frac{2^N}{\lambda^2}$  possible exceptions for all other sequences, i.e. for the large majority of all sequences,  $R(Q)$  is of order  $\frac{NQ^2}{8 \log Q} + O\left(\frac{NQ^2}{\log^2 Q}\right)$ .

Thus we have proved the following

THEOREM 1. Let us consider all  $2^N$  subsequences  $S_N$  of the sequence  $\{1, 2, \dots, \dots, N\}$ . We have for all these subsequences with the possible exception of  $\frac{2^N}{\lambda^2}$  such sequences

$$(2.22) \quad R(Q) = \frac{NQ^2}{8 \log Q} + O\left(\frac{NQ^2}{\log^2 Q}\right)$$

where  $Q \geq N^{1/3}$  and

$$(2.23) \quad \lambda = \min\left(\frac{Q}{(\log Q)^{3/2}}, \frac{\sqrt{N}}{\log Q}\right).$$

Thus, if  $Q \leq \sqrt{N \log N}$ , (2.22) holds except for at most  $\frac{2^N \log^3 Q}{Q^2}$  sequences, while for  $Q > \sqrt{N \log N}$  (2.22) holds, except for at most  $\frac{2^N \log^2 N}{N}$  sequences.

COROLLARY. If  $Q = \sqrt{AN \log N}$  ( $A > 1$ ) then  $R(Q) \sim \frac{AN^2}{8}$  except for at most  $\frac{2^N \log^2 N}{N}$  sequences.

Let us now consider the quantity

$$\text{Max}_{p \leq Q} \left[ \text{Max}_{0 \leq a \leq p-1} \left| Z(a, p) - \frac{Z}{p} \right| \right].$$

It is easy to show [using the central limit theorem and the fact that for any given  $p$  the quantities  $Z(a, p)$  ( $a = 0, 1, \dots, p-1$ ) are independent], that

$$(2.24) \quad P\left( \text{Max}_{0 \leq a \leq p-1} \left| Z(a, p) - \frac{Z}{p} \right| > \sqrt{\frac{N \log pQ}{2p}} \right) = O\left(\frac{1}{Q}\right)$$

and thus except for at most  $O\left(\frac{2^N}{\log Q}\right)$  exceptional sequences we have

$$\left| Z(a, p) - \frac{Z}{p} \right| \leq \sqrt{\frac{N \log pQ}{2p}}$$

for all  $a$  and  $p$  ( $0 \leq a \leq p-1$ ,  $p \leq Q$ ).

On the other hand, using again the independence of the random variables  $Z(a, p)$  ( $a = 0, 1, \dots, p-1$ ) and the central limit theorem it follows that for all except for at most  $O\left(\frac{2^N}{N}\right)$  sequences  $S^N$ , one has, for all  $p$  such that

$$C \log N < p < \frac{N}{\sqrt{\log N}}$$

$$\Delta^2(p) = \frac{Np}{4} \left( 1 + O\left(\frac{1}{\sqrt{\log N}}\right) \right)$$

if  $C$  is a sufficiently large positive number.

### § 3. The values of a random trigonometrical polynomial at well spaced points

In this § we shall consider the sum

$$(3.1) \quad T(\alpha, S) = \sum_{\nu=1}^D |S(\alpha_\nu)|^2$$

where  $\alpha_1, \alpha_2, \dots, \alpha_D$  are real numbers "well spaced" in the sense of Davenport and Halberstam, satisfying

$$(3.2) \quad 0 < \alpha_1 < \alpha_2 < \dots < \alpha_D < 1 \quad \text{and} \\ \alpha_{\nu+1} - \alpha_\nu \geq \delta > 0 \quad \text{for } \nu = 1, 2, \dots, D-1$$

and  $S(\alpha)$  is the random trigonometric polynomial

$$(3.3) \quad S(\alpha) = \sum_{n=1}^N \varepsilon_n e^{2\pi i n \alpha}$$

where  $\varepsilon_1, \dots, \varepsilon_N$  are independent random variables, each taking on the values 1 and 0 with probability  $\frac{1}{2}$ .

We first evaluate the expectation of  $T(\alpha, S)$ . We have clearly

$$(3.4) \quad T(\alpha, S) = \sum_{n=1}^N \sum_{m=1}^N \varepsilon_n \varepsilon_m \sum_{\nu=1}^D e^{2\pi i(n-m)\alpha_\nu}$$

and thus

$$(3.5) \quad E(T(\alpha, S)) = \frac{1}{2} \sum_{\nu=1}^D \left( \frac{N}{2} + \sum_{l=1}^{N-1} (N-l) \cos 2\pi l \alpha_\nu \right) + \frac{ND}{4}.$$

Now it is well known that

$$(3.6) \quad \frac{N}{2} + \sum_{l=1}^{N-1} (N-l) \cos 2\pi l \alpha_\nu = \frac{\sin^2 N\pi \alpha_\nu}{2 \sin^2 \pi \alpha_\nu}.$$

As a matter of fact the formula (3.6) is well known as a formula for Fejér's kernel of the arithmetic means of Fourier series.

It follows from (3.5) and (3.6) that

$$(3.7) \quad E(T(\alpha, S)) = \frac{1}{4} \sum_{\nu=1}^D \frac{\sin^2 N\pi \alpha_\nu}{\sin^2 \pi \alpha_\nu} + \frac{ND}{4}.$$

Let us now consider the special case when  $\alpha_Q^* = (\alpha_1^*, \dots, \alpha_D^*)$  is the set of all numbers  $\frac{a}{q}$  with  $(a, q) = 1$ ,  $1 \leq a \leq q$ ,  $1 < q \leq Q \leq N$ . It is easy to see that

$$(3.8) \quad \sum_{q \leq Q} \sum_{\substack{a=1 \\ (a,q)=1}}^q \frac{\sin^2 N\pi \frac{a}{q}}{\sin^2 \pi \frac{a}{q}} = O(Q^3)$$

thus, denoting by  $\varphi(q)$  the number of numbers  $a < q$  relatively prime to  $q$ , we have

$$(3.9) \quad E(T(\alpha_Q^*, S)) = \frac{N}{4} \sum_{q=1}^Q \varphi(q) + O(Q^3).$$

As however

$$(3.10) \quad \sum_{q=1}^Q \varphi(q) = \frac{3Q^2}{\pi^2} + O(Q \log Q)$$

it follows that

$$(3.11) \quad E(T(\alpha_Q^*, S)) = \frac{3Q^2N}{4\pi^2} + O(NQ \log Q) + O(Q^3).$$

It follows that for  $Q = o(N)$  there exists for each  $\varepsilon > 0$  a sequence  $S_N$  for which

$$(3.12) \quad T(\alpha_Q^*, S_N) > \frac{3Q^2N(1-\varepsilon)}{4\pi^2}.$$

Thus the estimate

$$(3.13) \quad \sum_{q \leq Q} \sum_{\substack{a=1 \\ (a,q)=1}}^q \left| S\left(\frac{a}{q}\right) \right|^2 = O(N^2)$$

which according to the theorem of Davenport and Halberstam is valid for  $Q \leq \sqrt{N}$  cannot be valid if  $Q$  is of larger order of magnitude than  $\sqrt{N}$ . By evaluating the variance of  $T(\alpha, S)$  one can prove even more, namely that  $T(\alpha_Q^*, S_N) \sim \frac{3Q^2N}{4\pi^2}$  for all except  $o(2^N)$  sequences  $S_N$ , if  $\frac{1}{Q} = o(1)$  and  $\frac{Q}{N} = o(1)$ . In particular one can prove that

$$(3.14) \quad D \left( \sum_{q \leq \sqrt{N}} \sum_{\substack{a=1 \\ (a,q)=1}}^{q-1} \left| S\left(\frac{a}{q}\right) \right|^2 \right) = O(N^{3/2})$$

which implies that except for at most  $O\left(\frac{2^N \log N}{N}\right)$  exceptional sequences

$$(3.15) \quad \sum_{q \leq \sqrt{N}} \sum_{\substack{a=1 \\ (a,q)=1}}^{q-1} \left| S\left(\frac{a}{q}\right) \right|^2 \sim \frac{3N^2}{4\pi^2}.$$

Let us summarize now our results: Theorem 1 shows that the estimate

$$(3.16) \quad R(Q) = O(N^2)$$

cannot hold if  $Q$  is of larger order of magnitude than  $\sqrt{N \log N}$ . It remains an open question whether (3.16) holds if  $\sqrt{N} \leq Q \leq \sqrt{N \log N}$ . However, (3.11) shows that even if (3.16) is true for the range  $\sqrt{N} \leq Q \leq \sqrt{N \log N}$  it cannot be proved by the methods used up to now, as all these methods gave estimates for  $R(Q)$  through estimating  $T(\alpha_Q^*, S)$ .

#### § 4. Some open problems

Let  $S_N$  denote a subsequence of the sequence  $\{1, 2, \dots, N\}$  which contains at least  $cN$  elements ( $0 < c < 1$ ). Let  $Y(\alpha, \varepsilon)$  where  $0 < \varepsilon < 1$  and  $1/2 \leq \alpha < 1$  denote the number of those primes  $p \leq N^\alpha$  for which at least  $p\varepsilon$  residue classes mod  $p$  do not contain any element of  $S_N$ . It follows already from Linnik's result (1.1) that  $Y\left(\frac{1}{2}, \varepsilon\right)$  is bounded, namely that

$$(4.1) \quad Y\left(\frac{1}{2}, \varepsilon\right) \leq \frac{20\pi}{c\varepsilon^2}.$$

From (1.12) one obtains the slightly better estimate

$$(4.2) \quad Y\left(\frac{1}{2}, \varepsilon\right) \leq \frac{\pi + 1}{\varepsilon c}.$$

As regards  $Y(\alpha, \varepsilon)$  with  $1/2 < \alpha < 1$  we get from (1.11) the estimate

$$(4.3) \quad Y(\alpha, \varepsilon) \leq \frac{N^{2\alpha-1}}{\varepsilon c} + \frac{\pi}{\varepsilon c}.$$

It seems probable that (4.3) is far from being best possible; it is an open problem whether  $Y(\alpha, \varepsilon)$  is bounded for every  $\alpha$  with  $\frac{1}{2} < \alpha < 1$ , or not. Of course,  $Y(1, \varepsilon)$  is not bounded: as a matter of fact if  $S_N$  is the sequence of numbers  $\leq Nc$  ( $0 < c < \frac{1}{2}$ ) and  $0 < \varepsilon < \frac{1}{2}$  then for all primes  $p$  with  $\frac{cN}{1-\varepsilon} < p < N$  at least  $p\varepsilon$  residue classes mod  $p$  do not contain any element of  $S_N$ , and thus

$$Y(1, \varepsilon) \geq \frac{N}{\log N} \left(1 - \frac{c}{1-\varepsilon}\right) + O\left(\frac{N}{\log^2 N}\right).$$

Another related problem is the following: if  $0 < \varepsilon < 1$  let  $S_N$  be a subsequence of the sequence  $\{1, 2, \dots, N\}$  such that for every  $p$  with  $A_\varepsilon < p < N^\alpha$  where  $A_\varepsilon > 0$ ,  $0 < \alpha < 1$  there are at least  $\varepsilon p$  residue classes mod  $p$  which do not contain any element of  $S_N$ . What is the maximum  $M_N(\varepsilon, \alpha)$  of the number of terms of such a sequence  $S_N$ ? It is easy to show that for each  $\varepsilon$  with  $0 < \varepsilon < 1/2$   $M_N(\varepsilon, 1) \geq [\sqrt{N}]$ . As a matter of fact let  $S_N$  denote the sequence of squares  $\leq N$ . Clearly if  $b$  is a quadratic non-residue mod  $p$ , then there is no element of the sequence  $1^2, 2^2, \dots, k^2, \dots$  which is congruent to  $b$  mod  $p$ ; thus for each  $p$  the number of empty residue classes is at least  $\frac{p-1}{2}$  if  $p \geq 3$ .

## References

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